

# ON APPROACH TO RANDOMNESS TESTING ON THE BASE OF THE UNIVERSAL PREDICTORS

A.L. KOSTEVICH, A.V. SHILKIN

*National Research Center for Applied Problems of Mathematics and Informatics  
Belarusian State University, Minsk, BELARUS*

e-mail: kostevich@bsu.by

## Abstract

The proposed approach allows to construct a statistical test for randomness using the known predictors that are universal for general alternative hypotheses. The test is shown to be consistent and to have a required significance level. An example of the test based on the Lempel-Ziv algorithm is given.

## 1 Introduction

Randomness testing is the topical problem in cryptography and simulation [5]. There are two traditional approaches to randomness testing [2]:

– **Empiric approach.** One proposes some statistic with a known probability distribution in the case of the random sequence and constructs a statistical test for randomness testing using this statistic. But alternative hypotheses detected by the constructed test and the power of the test are usually unknown.

– **Theoretical (parametric) approach.** One chooses some meaningful (parametric) alternative hypothesis and uses the statistical technique to construct an optimal statistical test for detection of the alternative hypothesis. But, on the one hand, a set of the alternative hypotheses (and number of the tests accordingly) in applications is usually large. On the other hand, construction of the single test for a general alternative hypothesis in most cases is impossible.

Therefore the alternative approach to randomness testing on the base of **the sequence complexity** is being developed [3, 4]. Loosely speaking, a sequence is considered to be random if it can not be compressed by a data compression method. There are a lot of universal data compression methods (e.g. the Lempel-Ziv algorithm), and they are widely used for randomness testing (see [4, 5]). However data compression methods usually have complicated algorithms and the investigation of their probabilistic properties is hard, e.g. both Maurer's test [4] and the Lempel-Ziv compression test [5] use estimates of unknown parameters of their statistics' distributions (see also section 3). Fortunately, B.Ya. Ryabko and V.A. Monarev succeeded in construction of the mathematically well-founded test for randomness testing on the base of an *arbitrary* universal data compression algorithm [6].

In the paper we propose an alternative approach to randomness testing on the base of an *arbitrary* universal predictor. We were motivated by Yao's theorem [1] about universality of the next-bit test and close relation between universal predictors and both classical statistical methods and data compression methods.

## 2 Randomness testing using a universal predictor

Let  $X_1, X_2, \dots$  be a sequence of binary random variables ( $X_t \in \mathcal{A} = \{0, 1\}$ ) described by a set of conditional probabilities  $\{\mathbf{P}\{X_{t+1} \mid X_t, \dots, X_1\}\}$ . One observes the first  $t$  random variables and tries to predict the next outcome  $X_{t+1}$ .

If the probability model of the sequence is known, then the optimal forecast  $X_{t+1}^*$  is determined by the maximum of the corresponding conditional probability and the minimum prediction error probability  $\pi_t^*$  is attained:

$$\begin{aligned} X_{t+1}^* &= \arg \max_{a \in \mathcal{A}} \mathbf{P}\{a \mid X_t, X_{t-1}, \dots, X_1\}, \\ \pi_t^*(X_t, \dots, X_1) &= \mathbf{P}\{X_{t+1}^* \neq X_{t+1} \mid X_t, \dots, X_1\} = \min_{a \in \mathcal{A}} \mathbf{P}\{a \mid X_t, X_{t-1}, \dots, X_1\}. \end{aligned} \quad (1)$$

If the probability model of the sequence is unknown, then the ‘‘prediction’’ conditional probabilities  $\{\hat{\mathbf{P}}\{X_{t+1} \mid X_t, \dots, X_1\}\}$  should be defined. After that, the next outcome can be predicted as in (1) with the greater prediction error probability  $\hat{\pi}_t$ :

$$\begin{aligned} \hat{X}_{t+1} &= \arg \max_{a \in \mathcal{A}} \hat{\mathbf{P}}\{a \mid X_t, X_{t-1}, \dots, X_1\}, \\ \hat{\pi}_t(X_t, \dots, X_1) &= \mathbf{P}\{\hat{X}_{t+1} \neq X_{t+1} \mid X_t, \dots, X_1\} \geq \pi_t^*(X_t, \dots, X_1). \end{aligned} \quad (2)$$

**Definition** ([7]). *A predictor (2) is called universal for the class  $\mathcal{M}$  if the prediction error tends to zero:  $\hat{\pi}_t(X_t, \dots, X_1) - \pi_t^*(X_t, \dots, X_1) \rightarrow 0$  in probability as  $t \rightarrow \infty$  w.r.t. any set of the conditional probabilities  $\{\mathbf{P}\{X_{t+1} \mid X_t, \dots, X_1\}\}$  from the class  $\mathcal{M}$ .*

It should be noted that there are other definitions of the universal predictor (see [7]) that differ from the given definition in specification of the prediction error. One can use the divergence or logarithm of the likelihood ratio for the conditional probabilities  $\{\mathbf{P}\{X_{t+1} \mid X_t, \dots, X_1\}\}$  and  $\{\hat{\mathbf{P}}\{X_{t+1} \mid X_t, \dots, X_1\}\}$  as the prediction error. In the last case and under the *parametric* description of the class  $\mathcal{M}$  the prediction error coincides with the well-known likelihood ratio statistic.

We shall use  $\hat{\pi}_t - \pi_t^*$  as the error probability of the predictor. This choice is more suitable for construction of a statistical test for randomness against an alternative hypothesis described by a general (*nonparametric*) class  $\mathcal{M}$ .

Let us define an auxiliary sequence of indicators of successful predictions. Let  $n$  binary random variables  $X_1, \dots, X_n$  be observed. For each  $t = 1, 2, \dots, n-1$  we will predict the next outcome  $X_{t+1}$  using the predictor (2), which is constructed on the base of the first  $t$  r.v.  $X_1, \dots, X_t$ , and  $\hat{X}_1 = 1$  for simplicity. After that, for each  $t = 1, 2, \dots, n$  we calculate the indicator the successful prediction:  $Y_t = \mathbf{1}\{\hat{X}_t = X_t\}$ . It is easily seen that the sequence  $\{Y_t\}$  has the following properties:

**P1.** If the predictor (2) is deterministic, then the sequence  $\{X_t\}$  can be restored from the sequence  $\{Y_t\}$  (i.e. there is no loss of information when considering only  $\{Y_t\}$ ), so  $\mathbf{P}\{Y_t = 1 \mid Y_{t-1}, \dots, Y_1\} = \mathbf{P}\{Y_t = 1 \mid X_{t-1}, \dots, X_1\}$ .

**P2.** If the predictor (2) is universal, then as  $t \rightarrow \infty$

$$\mathbf{P}\{Y_t = 1 \mid X_{t-1}, \dots, X_1\} - \max_{a \in \mathcal{A}} \mathbf{P}\{a \mid X_{t-1}, \dots, X_1\} \rightarrow 0 \text{ in probability.}$$

Let us consider a null hypothesis  $\mathcal{H}_0$  that the sequence  $\{X_t\}$  is random:

$$\mathcal{H}_0 : \quad \{X_t\} \text{ are i.i.d. Bernoulli r.v., } \mathbf{P}\{X_t = 0\} = \mathbf{P}\{X_t = 1\} = \frac{1}{2}.$$

Clear, under  $\mathcal{H}_0$  the indicators  $\{Y_t\}$  are also i.i.d. Bernoulli r.v. with  $\mathbf{P}\{Y_t = 1\} = 0.5$ .

Let us consider an alternative hypothesis  $\mathcal{H}_1$  that the sequence  $\{X_t\}$  is described by the following conditional probabilities:

$$\mathcal{H}_1 : \quad \max_{a \in \mathcal{A}} \mathbf{P}\{a \mid X_{t-1}, \dots, X_1\} = \frac{1}{2} + \varepsilon_{t, X_{t-1}, \dots, X_1}, \quad \varepsilon_{t, x_{t-1}, \dots, x_1} \geq 0. \quad (3)$$

If the predictor (2) is universal for  $\mathcal{H}_1$ , then the probability distribution (3) induces an unknown multivariate probability distribution of  $\{Y_t\}$  with the marginal probabilities:

$$\mathcal{H}_1^{(Y)} : \quad P\{Y_t = 1\} = \frac{1}{2} + \varepsilon_t, \quad \varepsilon_t \geq 0 \quad \text{as } t \rightarrow \infty.$$

In this case the natural statistical test for randomness is based on the frequency of the successful predictions and has the form:

$$\text{decide } \begin{cases} \mathcal{H}_0, & \text{if } 2\sqrt{n}(S - \frac{1}{2}) < \Phi^{-1}(1 - \alpha), \\ \mathcal{H}_1, & \text{otherwise,} \end{cases} \quad S = \frac{1}{n} \sum_{t=1}^n Y_t, \quad (4)$$

where  $\Phi(\cdot)$  is the standard normal c.d.f.,  $\alpha$  is a significance level.

**Claim 1.** *If the test (4) is based on an arbitrary predictor, then the type I error probability of the test tends to  $\alpha$  as  $n \rightarrow \infty$ .*

**Claim 2.** *If the test (4) is based on the universal for  $\mathcal{H}_1$  predictor,  $\{Y_t\}$  satisfy the central limit theorem and  $\mathbf{D}\{S\} \asymp n^{-1}$  as  $n \rightarrow \infty$ , then the power of the test is:*

$$W_n \approx \Phi\left(c_1 \frac{1}{\sqrt{n}} \sum_{t=1}^n \varepsilon_t - c_2\right) \quad \text{and} \quad W_n \rightarrow 1 \quad \text{as} \quad \frac{1}{\sqrt{n}} \sum_{t=1}^n \varepsilon_t \rightarrow +\infty,$$

where  $c_1, c_2$  are positive constants.

### 3 Test based on the Lempel-Ziv predictor

It is known that the Lempel-Ziv predictor is universal for the class  $\mathcal{M}$  of stationary and ergodic Markov sources of finite order. The Lempel-Ziv algorithm partitions a sequence into phrases of variable size such that a new phrase is the shortest subsequence not seen in the past as a phrase. The Lempel-Ziv algorithm predicts the next outcome (see (2)) as the most frequent suffix given a prefix of a phrase.

The Lempel-Ziv test of NIST [5] is based on the number of phrases in a sequence of length  $n = 10^6$ . But it is noted in [5] that the conditions of application of theoretical mean 50171.7 and variance 33.59 of this statistic remain ambiguous, and the statistical estimates 69586.25 and 70.44 of that parameters are taken.

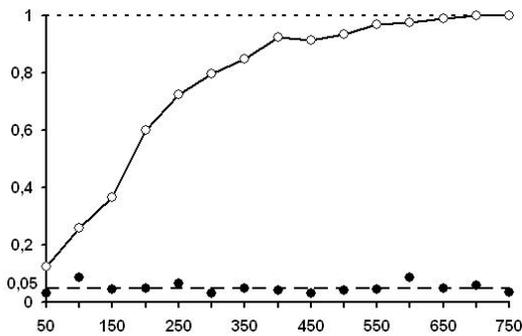


Figure 1: Performance of the test

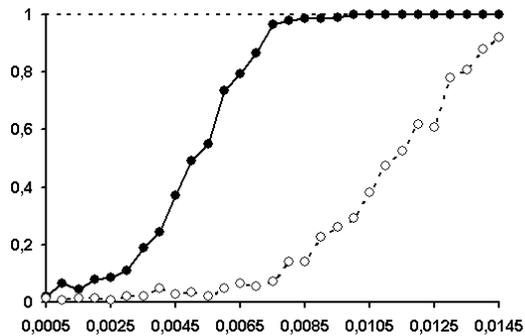


Figure 2: Comparison of the powers

In contrast to [5] and other approaches, we proposed the well-founded way to construct the test on the base of the Lempel-Ziv algorithm or any other predictor.

We conducted simulation experiments to estimate the performance of the test (4) based on the Lempel-Ziv predictor. As the hypothesis  $\mathcal{H}_1$  we used the Markov chain model with  $\mathbf{P}\{X_{t+1} = 1 \mid X_t = 0\} = \mathbf{P}\{X_{t+1} = 0 \mid X_t = 1\} = 0.5 + \varepsilon$ . Figure 1 presents calculated estimates of the type I error probability (denoted by  $\bullet$ ) and the power (denoted by  $\circ$ ) of the proposed test w.r.t. length  $n$ , where  $\alpha = 0.05$ ,  $\varepsilon = 0.15$ . One can see that the proposed test is consistent and the type I error probability of the test is approximately  $\alpha$ . Figure 2 presents calculated estimates of the power of the proposed test (denoted by  $\bullet$ ) and the Lempel-Ziv test of NIST [5] (denoted by  $\circ$ ) w.r.t the parameter  $\varepsilon$  of the Markov chain of length  $n = 10^6$ . One can see that the power of the proposed test is greater than the power of NIST's test.

## References

- [1] Goldreich O., Goldwasser S., Micali S. (1986). How to construct random functions. *J. of the Association for Computing Machinery*. Vol. **33**, pp. 792-807.
- [2] Knuth E.E. (1981). *The art of computer programming*, vol. 2. Addison-Wesley.
- [3] Kolmogorov A.N. (1965). Three approaches to the quantitative definition of information. *Problems of Information Transmission*. Vol. **1**, pp. 3-11.
- [4] Maurer U. (1992). A universal statistical test for random bit generators. *J. of Cryptology*. Vol. **5** (2), pp. 89-105.
- [5] NIST Special Publication 800-22. (2001). *A statistical test suite for random and pseudorandom number generators for cryptographic applications*.
- [6] Ryabko B.Ya., Monarev V.A. (2005). Using information theory approach to randomness testing. *J. of Statistical Planning and Inference*. Vol. **133** (1), pp. 95-110.
- [7] Suzuki J. (2003). Universal prediction and universal coding. *Systems and Computers*, Vol. **34** (6), pp. 1-11.